2025 May 20 : CIAO Workshop UMass Lowell

### Source Detection & Statistics

#### VINAY KASHYAP CfA/CXC-Calibration

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### Basics of high-energy Astrostatistics

#### VINAY KASHYAP CfA/CXC-Calibration



- \* Why bother?
- \* What to bother about
  - \* errors, uncertainties, and appropriateness
  - \* detection, fitting :: deciding, estimating
- \* Where CIAO/Sherpa is doing the bothering for you

### Outline

In the space of one hundred and seventy-six years the Lower Mississippi has shortened itself two hundred and forty-two miles. Therefore, any calm person, who is not blind or idiotic ... can see that seven hundred and forty-two years from now the lower Mississippi will be only a mile and three-quarters long.

There is something fascinating about science. One gets such wholesale returns of conjecture out of such a trifling investment of fact.

-Mark Twain, Life on the Mississippi

## Lessons from Sam Clemens

- \* Generalization is the foundational cornerstone of science, especially of astronomy, and it is trivially easy to shoot yourself in the foot
- \* It is easy to run into situations where fluctuations can ruin an observation, or you push your model of the Universe too far into realms where it is not applicable
- We must pay attention to errors and uncertainties in the data, on the quality of the models we build to explain them, and how we make decisions and draw inferences



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- \* Characterizing a model from measurements requires a great deal of logical and rigorous thinking

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- Interpretation of observations needs a model
- \* Characterizing a model from measurements requires a great deal of logical and rigorous thinking
- Astrostatistics is a system to
  - quantify uncertainty
  - extend inference to complex systems
  - make decisions



## Error bars



- Invented by Gauss while figuring out where Ceres would reappear after coming around the Sun: \*
  - Measurement uncertainties follow a "normal" distribution, with deviations being symmetrical, large \* discrepancies being less frequent, and the average represents the likely value



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  - Measurement uncertainties follow a "normal" distribution, with deviations being symmetrical, large discrepancies being less frequent, and the average represents the likely value
  - \* Lets you fit a simplified model curve to the data by minimizing the squared deviations of residuals





# Error propagation : Gaussian

- Simple case: if everything is distributed as a Gaussian, and has well-defined means and standard \* deviations, then at "best fit" values  $a_i$ ,  $g = g(a_i)$

$$\sigma_g^2 = \sum_i \frac{1}{N} \sum_k (g_k(a_i + \delta a_i) - g_k(a_i))^2 \text{ for all data p}$$

and expand as Taylor series and sum over k to get to the  $2^{nd}$  order

$$\sigma_g^2 = \sum_i \sum_j \frac{\partial g}{\partial a_i} \frac{\partial g}{\partial a_j} \sigma_{a_i a_j}$$
 where *i*, *j* are variables

or ignoring correlations amongst the  $\{a_i\}, \sigma_{a_i a_i} = \sigma_{a_i}^2 \delta_{ij}$ 

$$\sigma_g^2 \approx \sum_i \left(\frac{\partial g}{\partial a_i}\right)^2 \sigma_{a_i}^2$$

\* How to propagate uncertainty from one stage to another — if g = f(x), and  $\sigma_x$  is known, what is  $\sigma_g = f(\sigma_x)$ 

points k=1..N and independent variables i



 $g = g(a_i)$  $\sigma_g^2 = \sum_{i} \left(\frac{\partial g}{\partial a_i}\right)^2 \sigma_{a_i}^2$ 

uncertainties scale (counts  $\rightarrow$  count rate)

 $g = g(a_i)$  $\sigma_g^2 = \sum_i \left(\frac{\partial g}{\partial a_i}\right)^2 \sigma_{a_i}^2$ 

 $g = C \cdot a \rightarrow \sigma_g = C \cdot \sigma_a$ 

- $g = \ln(a)$ 
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  - fraction

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  - g = a + berrors s

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$$\rightarrow \sigma_g^2 = \sigma_a^2 + \sigma_b^2$$
quare-add
9

0.1%

μ-3σ

- \* Equal tail
- \* Highest density (also the shortest interval)

\* "
$$1\sigma$$
"  $\Rightarrow 68\% \equiv$ 

16<sup>th</sup>-84<sup>th</sup> percentile

- \*  $5^{\text{th}}-95^{\text{th}} \equiv 90\% \equiv 1.6\sigma$
- \*  $2.5^{\text{th}} 97.5^{\text{th}} \equiv 95\% \equiv 2\sigma$

\* " $3\sigma$ "  $\Rightarrow p=0.003$ 







A distribution is described by a mathematical function, but is not simply a locus of points. It describes probabilities of sampling, of possible results of repeated experiments. Also describes the *likelihood* of seeing the observed data given the particular model parameter values



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Gaussian: 
$$N(x; \mu, \sigma) = -\frac{1}{\sigma \sqrt{2}}$$
  
Poisson:  $P(X = k; \lambda) = -\frac{1}{\beta}$   
Gamma:  $\gamma(x; \alpha, \beta) = -\frac{1}{\Gamma(\alpha)}$ 





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Poisson: $P(X = k; \lambda) = -\frac{1}{k}$ Gamma: $\gamma(x; \alpha, \beta) = \frac{1}{\Gamma(\alpha)}$ Delta: $\delta(x - x_0) = 0$  fo

 $\{X\} \sim f(\theta)$  $\int d\theta f(\theta) = 1$ 



or  $x \neq x_0$ , and  $\int dx \, \delta(x - x_0) = 1$ 



- \* *p*-value: the tail integral of a distribution
- \* False +ve  $\equiv$  Type I error: the *p*-value at which a threshold is set for a *null* distribution
- \* False –ve = Type II error: the fraction of the alternate distribution that falls below the threshold

0.30  $p(S|\lambda_{s}=0,\lambda_{B},r,\tau_{s},\tau_{B})$ 0.25 0.20 0.15 0.10 0.05 0.15  $p(S|\lambda_{S},\lambda_{B},r,\tau_{S},\tau_{B})$ 0.10 0.05



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## that the risk of false +ve is deemed acceptable

#### Orion Trapezium Star Forming Region

10 arcsec

An observation that falls in the tail of the null distribution where the *p*-value is small enough



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 $\Rightarrow$  95±11 counts from source

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*Estimate* source strength as  $Net = N_S - N_B / B$  $\sigma_{Net} = \sqrt{\{N_S + N_B / \mathbb{B}^2\}}$ 

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1 - 0.842 = 0.158



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## Fitting: Best-fit

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- \* e.g., linear regression  $y_i = \alpha + \beta x_i + \epsilon$ minimizing sum-squared residuals)

$$\ln L \propto \sum_{k} (y_{k} - \alpha - \beta x_{k})^{2}, \text{ with } \frac{\partial \ln L}{\partial \alpha} = \frac{\partial \ln L}{\partial \beta}$$
$$\hat{\beta} = \operatorname{Cov}(x, y) / \operatorname{Var}(x) \equiv \rho(x, y) \sqrt{\frac{\operatorname{Var}(x)}{\operatorname{Var}(y)}}, \text{ and } \hat{\alpha} =$$

#### solve by finding extremum of log likelihood (for Gaussian case, maximizing likelihood means

= 0

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Notice notation:

 $\overline{bar}$  and  $\hat{h}\hat{a}\hat{t}$  to indicate sample averages and best-fit values Грєєк letters for model quantities, Roman for data quantities

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# Fitting: Error Bars

\* Covariance errors aka curvature errors aka inverse of the Hessian

For Gaussian,  $\frac{\partial^2 \ln L_{\text{Gauss}}}{\partial x^2} \propto \frac{1}{\sigma^2}$ 

- i.e., compute curvature of log-likelihood surface at best fit and return its inverse as the variance
- + easy, obtained as byproduct of fitting
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### Fitting: Error Bars

#### \* $\Delta \chi^2$

Difference from best-fit  $\chi^2$  value is itself a  $\chi^2$  distribution with dof=1, so look for percentiles of that distribution:

$$\Delta \chi^2 = +1 \equiv 68\% \ (1\sigma)$$

 $\Delta \chi^2 = +2.71 \equiv 90\% (1.6\sigma)$ 

+ more robust than curvature

– gets complicated if parameters are correlated

- higher computational cost



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- \* How can you tell when you *do* have a "good" fit?

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\*  $-2 \ln L_{Gauss}$  is called the chi-square,

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- \* and its distribution describes the probability of getting  $(x_k, y_k)$  to match "similarly" for several bins
- \* When the observed  $\chi^2 \sim dof \pm \sqrt{2} \cdot dof$ , the model is doing a good job of matching the data. The farther it is from this range, the less likely it is that the model is a good description of the data
  - But always use your judgement, because this is a *probabilistic* rule! \*
  - \* Watch out for how  $\sigma^2$  is defined (model variance is better)

## Fitting: CStat

\* Recall Poisson log Likelihood:  $\ln L_{\text{Pois}} = -\ln \Gamma(k+1) + k \ln \lambda - \lambda$ \*  $\chi^2$  is  $-2 \ln L_{\text{Gauss}}$ , cstat is  $-2 \ln \frac{L_{\text{Pois}}}{L_{\text{Saturated Pois}}}$ 

 $\mathbf{cstat} = 2\sum_{i} M_i - D_i + D_i \cdot (\ln D_i - \ln M_i)$ 

- \* Watch out: cstat is only asymptotically  $\chi^2$ , not quite the Poisson likelihood, 0s are thrown away, background must be explicitly modeled
- \* unbiased for low counts compared to  $\chi^2$ , asymptotically  $\chi^2$ , rudimentary goodness-of-fit exists (Kaastra 2017, A&A 605, A51) [AnetaS] https://cxc.cfa.harvard.edu/ciao/workshop/oct20\_egypt\_virt/cstat\_vs\_chisq\_SimsNotebook.ipynb [AnetaS] https://cxc.cfa.harvard.edu/ciao/workshop/oct20\_egypt\_virt/data\_for\_cstat\_vs\_chisq\_SimsNotebook.tar.gz

where  $D_i$  are observed counts, and  $M_i$  are model predicted counts in bin *i* and Saturated Poisson means  $\lambda$  is set to k, or a model where  $M_i$  is set to  $D_i$ . cstat can also be derived as an approximate form of the  $L_{Pois}$ , using Stirling's approximation for  $\Gamma(\cdot)$ 

Handbook of X-ray Astronomy (Arnaud, Smith, Siemiginowska): https://doi.org/10.1017/CBO9781139034234.008



photon index is marked with a dashed line and it was set at  $\gamma = 1.28$ 

Fig. 7.3 Distributions of a photon index parameter  $\gamma$  obtained by fitting simulated X-ray spectra with 6000 counts and using the three different statistics:  $S_{\text{Pearson}}^2$ ,  $S^2$  and C (i.e. the Poisson likelihood) statistics. The true value of the simulated

# Statistical Tools in CIAO/Sherpa

- \* fit: non-linear minimization fitting
- \* **conf/covar**: uncertainty intervals and error bars
- \* resample\_data: to get bootstrap distribution of model parameter draws when data errors are asymmetric
- \* **bootstrap/sample\_flux/sample\_photon\_flux/sample\_energy\_flux**: with replacement/parametric bootstrap to get Monte Carlo distribution accounting for parameter uncertainties
- \* **get\_draws**: Markov Chain Monte Carlo (MCMC) engine **pyBLoCXS** (Bayesian Low-Counts X-ray Spectral analysis; van Dyk et al. 2001, ApJ 548, 224)
- \* calc\_mlr, calc\_ftest: model comparison via Likelihood Ratio Test (LRT)/F-test
- \* plot\_pvalue, plot\_pvalue\_results: to do posterior predictive p-value checks (Protassov et al. 2002, ApJ 571, 545)
- \* glvary: light curve modeling (Gregory & Loredo 1992, ApJ 398, 146)
- \* celldetect/wavdetect/vtpdetect/mkvtpbkg: source detection in images
- \* aprates: Bayesian aperture photometry also used in **srcflux** (Primini & Kashyap 2014, ApJ 796, 24)
- \* the python interpreter in Sherpa gives access to python libraries, and can be used to call upon packages and libraries in R, which are written by statisticians for statisticians

### Statistics is a tool; it can be misused

- \* "All models are wrong, but some are useful." George Box, c.1987
- Tak, H., et al., 2024, ApJS 275, 30 <u>https://doi.org/10.3847/1538-4365/ad8440</u> —
   be aware of
  - \* where the data are coming from, what sort of filtering they have been subjected to
  - \* conditions under which a particular theorem is proven to be valid and make sure they apply under the scenario you are considering
  - \* assumptions made about the nature and range of model parameters and model complexity, and always do sanity checks and sensitivity analyses
  - limitations of hypothesis tests, and always consider the power of the tests and false detection rates



# Some useful reading material

- \* Larry Bretthorst (1988), Bayesian Fourier analysis, https://bayes.wustl.edu/glb/book.pdf
- p/book/9780412983917
- \* Larry Wasserman (2006), All of Non-Parametric Statistics, <u>http://www.stat.cmu.edu/~larry/all-of-nonpar/</u>
- \* Rasmussen & Williams (2006), Gaussian Processes for Machine Learning, http://www.gaussianprocess.org/gpml/
- \* Feigelson & Babu (2012), Modern Statistical Methods for Astronomy with R Applications, https://astrostatistics.psu.edu/MSMA/
- analysis-for-the-physical-sciences/09E9A95DAE275F5B005676C71B542598
- \* Gelman et al. (2013), Bayesian Data Analysis, https://www.routledge.com/Astrostatistics-1st-Edition/Babu-Feigelson-Morgan-Keiding-Van-der-Heijden/p/book/9780412983917
- \* Edward Robinson (2016), Data analysis for scientists and engineers, <u>https://press.princeton.edu/titles/10911.html</u>
- \* Vinay Kashyap (2020), Basics of Astrostatistics, https://iachec.org/wp-content/uploads/2021/05/Kashyap\_2020\_Tutorial\_Guide\_to\_X-
- \* Buchner & Boorman (2023), Statistical aspects of X-ray spectroscopy, https://arxiv.org/abs/2309.05705
- ad8440

\* Tom Loredo (1990), monograph on neutrinos from 87A, http://hosting.astro.cornell.edu/staff/loredo/bayes/L90-LaplaceToSN1987A-scan.pdf

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\* Arnaud, Smith, & Siemiginowska (2011), Handbook of X-ray Astronomy, http://hea-www.cfa.harvard.edu/~rsmith/xrayastronomyhandbook/

\* Phil Gregory (2012), Bayesian Logical Data Analysis for Physical Sciences, https://www.cambridge.org/core/books/bayesian-logical-data-

ray\_and\_Gamma-ray\_Astronomy\_Data\_Reduction\_and\_Analysis\_Editor\_Cosimo\_Bambi\_Springer\_ISBN\_978-981-15-6337-9\_chapter\_6-1.pdf \* Feigelson, Kashyap, Siemiginowska (2022), Time domain methods for X-ray and gamma-ray astronomy, https://arxiv.org/abs/2203.08996 \* Tak et al. (2024), Six Maxims of Statistical Acumen for Astronomical Data Analysis, ApJS 275, 30 https://doi.org/10.3847/1538-4365/







